# Einstein's derivation of Brownian transition density

Particle starts at  $0 \in \mathbf{R}^3$  and is pushed around by tiny molecular bombardments.

 $f(t, x)dx = P(X_t \in dx) = \lim_{h \to 0} h^{-3}P(X_t \in a \text{ box of side length } h \text{ around } x).$ 

$$p(s, x, t, y)dy = P(X_t \in dy \mid X_s = x).$$

$$f(t+\tau,x) = \int f(x-y,t) p(x-y,t,x,t+\tau) dy.$$

homogeneity in space and time p(s, x, t, y) = p(t - s, y - x).

$$f(t + \tau, x) = \int (f(t, x) - y \cdot \nabla f(t, x) + \frac{1}{2}y \cdot D^2 f(t, x)y + \cdots)p(\tau, y)dy$$
  
$$= f(t, x) \int p(\tau, y)dy - \sum_{i=1}^3 \frac{\partial f}{\partial x_i}(t, x) \int y_i p(\tau, y)dy$$
  
$$+ \frac{1}{2} \sum_{i,j=1}^3 \frac{\partial^2 f}{\partial x_i \partial x_j}(t, x) \int y_i y_j p(\tau, y)dy + \cdots$$
  
$$\int p(\tau, y)dy = 1.$$

symmetry  $\int y_i p(\tau, y) dy = 0$  and  $\int y_i y_j p(\tau, y) dy = 0$   $i \neq j$ .

Influence of molecular bombardment in any two nonoverlapping intervals of time is independent

Variance should grow linearly, like the sum of independent random variables

$$Var(X_1 + \cdots + X_N) \simeq CN$$

$$\int y_i^2 p(\tau, y) dy = D\tau.$$

Letting au 
ightarrow 0 get heat equation

$$\frac{\partial f}{\partial t} = \frac{1}{2} D \Delta f.$$

With the obvious initial condition  $f(x, 0) = \delta_0$  this has the well known solution

$$f(t,x) = \frac{e^{-\frac{|x|^2}{2Dt}}}{(2\pi Dt)^{3/2}}.$$

# Markov processes

A process  $X_t$ ,  $t \ge 0$  is called a Markov process if for any function g and any  $t \ge s$ ,

$$\mathsf{E}[g(X_t) \mid X_u, \ 0 \le u \le s] = \mathsf{E}[g(X_t) \mid X_s].$$

Process determined by initial distr  $P(X_0 \in A)$  and the transition probs

$$p(s, x, t, A) = P(X_t \in A \mid X_s = x) \qquad s < t$$

$$P(X_{t_1} \in A_1, \dots, X_{t_n} \in A_n) = \int_{A_{n-1}} \cdots \int_{A_1} \int P(X_0 \in dx_0) p(0, x_0, t_1, dx_1) \cdots p(t_{n-1}, x_{n-1}, t_n, A_n)$$

Chapman-Kolmogorov equations

$$p(s, x, t, A) = \int p(s, x, u, dy) p(u, y, t, A)$$
 for  $s \le u \le t$ .

Example. Brownian motion  $p(s, x, t, dy) = \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{(y-x)^2}{2(t-s)}} dy$ 

# Gaussian measures

### Definition

Let  $(E, \mathcal{E}, \mu)$  be  $\sigma$ -finite, separable. There exists a Gaussian family  $\{X(f)\}_{f \in L^2(E, \mathcal{E}, \mu)}$  satisfying

•  $f \mapsto X(f)$  is linear

**2** 
$$E[X(f)] = 0, E[|X(f)|^2] = ||f||^2_{L^2(E,\mathcal{E},\mu)}$$

We can write  $X(A) = X(1_A)$ ,  $A \in \mathcal{E}$  to get a random measure

### Example: Brownian motion

 $X_t - X_s = X([s, t))$  is a Gaussian measure intensity  $\mu$  = Lebesgue measure

$$X(f) = \int f dB$$

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# Gaussian measures

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### Proof.

Let  $\{e_n\}_{n=1,2,3...}$  be an orthonormal basis of  $L^2(E, \mathcal{E}, \mu)$  and  $X_n$ , n = 1, 2, ... be iid  $\mathcal{N}(0, 1)$ .

$$X(f) = \sum_{n=1}^{\infty} \langle f, \boldsymbol{e}_n \rangle X_n$$

# Functions of finite variation

 $f: \mathbb{R}_+ \to \mathbb{R}$  right continuous

 $\Delta = 0 = t_0 < t_1 < \cdots < t_n = t \text{ subdivision of } [0, t]$ 

 $|\Delta| = \sup_i |t_{i+1} - t_i| = \text{mesh size}$ 

f is of finite variation if for each  $t < \infty$ ,  $\|f\|_{TV,[0,t]} = \sup_{\Delta} \sum_{i} |f(t_{i+1}) - f(t_{i})| < \infty$ 

## Proposition

- A function of finite variation is the difference of two monotone increasing functions
- μ([0, t]) = f(t) provides a 1-1 correspondence between measures on R<sub>+</sub> and functions of finite variation
- 3  $\int_0^\infty g(t)df(t) = \int gd\mu$  is the Riemann-Stieltjes integral
- A function of finite variation is differentiable almost everywhere

#### Definition

A stochastic process  $X_t$ ,  $t \ge 0$  has finite quadratic variation if there exists a finite process  $\langle X, X \rangle_t$ ,  $t \ge 0$ s.t. for each  $t < \infty$  and each sequence  $\{\Delta_n\}_{n=1,2,...}$  of subdivisions of [0, t] with  $|\Delta_n| \to 0$ ,

$$\lim_{n\to\infty}\sum_{i}|X_{t_{i+1}}-X_{t_i}|^2\stackrel{\text{prob}}{=}\langle X,X\rangle_t$$

The process  $\langle X, X \rangle_t$ ,  $t \ge 0$  is non-decreasing. It is called the *quadratic variation* of *X*.

Recall  $\lim_{n\to\infty} X_n \stackrel{\text{prob}}{=} X$  if  $P(|X_n - X| \ge \epsilon) \to 0$  for each  $\epsilon > 0$ 

Theorem

Let X be a Gaussian measure with intensity  $\mu$  on  $(E, \mathcal{E})$  Let  $A \in \mathcal{E}$  with  $\mu(A) < \infty$  Let  $\{A_k^n\}_{n=1,2,...}$  be finite partitions of A such that  $\sup_k \mu(A_k^n) \to 0$  as  $n \to \infty$  Then

$$\lim_{n\to\infty}\sum_k |X(A_k^n)|^2 = \mu(A)$$

in  $L^2(\Omega, \mathcal{F}, P)$ 

Proof.

 $\{X(A_{k}^{n})\}_{k}$  independent  $\mathcal{N}(0, |\mu(A_{k}^{n})|)$  so

$$\begin{aligned} E[|\sum_{k} X^{2}(A_{K}^{n}) - \mu(A_{k}^{n})|^{2}] &= \sum_{k} E[|X^{2}(A_{K}^{n}) - \mu(A_{k}^{n})|^{2}] \\ &= 2\sum_{k} |\mu(A_{k}^{n})|^{2} \leq 2\mu(A) \sup_{k} \mu(A_{k}^{n}) \end{aligned}$$

Theorem

Let X be a Gaussian measure with intensity  $\mu$  on  $(E, \mathcal{E})$  Let  $A \in \mathcal{E}$  with  $\mu(A) < \infty$ Let  $\{A_k^n\}_{n=1,2,\dots}$  be finite partitions of A such that  $\sup_k \mu(A_k^n) \to 0$  as  $n \to \infty$ Then

$$\lim_{n\to\infty}\sum_k |X(A_k^n)|^2 = \mu(A)$$

in  $L^2(\Omega, \mathcal{F}, P)$ 

### **Brownian motion**

Brownian motion is Gaussian measure with intensity dt

$$\langle B, B \rangle_t = t$$

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Note that if  $X_n \stackrel{prob}{\to} X$  one can always choose a (non-random) subsequence such that  $X_n \stackrel{prob}{\to} X$ So one can choose partitions so that

$$\lim_{n\to\infty}\sum_{i}|X_{t_{i+1}}-X_{t_i}|^2\stackrel{a.s.}{=}\langle X,X\rangle_t$$

For Brownian motion it turns out that any sequence  $\Delta_n \subset \Delta_{n+1}$  gives a.s. convergence

## Proposition

With probability one Brownian motion  $B_t$ ,  $t \ge 0$  is not of finite variation in any interval

#### Proof.

Let f be any continuous function on [0, t]

$$\sum_{i} |f_{t_{i+1}} - f_{t_i}|^2 \le \max_{i} |f_{t_{i+1}} - f_{t_i}| \sum_{i} |f_{t_{i+1}} - f_{t_i}|$$

Since  $\max_i |f_{t_{i+1}} - f_{t_i}| \rightarrow 0$ , if

$$\langle f, f \rangle_t = \lim \sum_i |f_{t_{i+1}} - f_{t_i}|^2 > 0$$

then

$$\lim \sum_{i} |f_{t_{i+1}} - f_{t_i}| = \|f\|_{TV,[0,t]} = \infty$$

and if  $||f||_{TV,[0,t]} < \infty$  then  $\langle f, f \rangle_t = 0$ 

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#### Proposition

With probability one Brownian motion  $B_t$ ,  $t \ge 0$  is not locally Hölder of order  $\alpha$  for any  $\alpha > 1/2$ 

#### Proof.

Let *f* be any continuous function on [0, t] s.t. for some  $0 \le a < b \le t$ and some  $\alpha > 1/2$ , for all  $a \le s, t \le b$ ,

$$|f_t - f_s| \le k |t - s|^{\alpha}$$

Then

$$\sum_{i} |f_{t_{i+1}} - f_{t_i}|^2 \le k^2 (b-a) \max_{i} |t_{i+1} - t_i|^{2\alpha - 1}$$

## Theorem. (Paley, Wiener, Zygmund 33)

Brownian motion is nowhere differentiable with probability one

Proof. (Dvoretsky, Erdös, Kakutani 61) Suppose that B(t) was differentiable at a point  $s \in [0, 1]$ .

Then  $\exists \epsilon > 0$  and an integer  $\ell \ge 1$  such that

 $|B(t) - B(s)| \le \ell(t-s)$  for  $0 < t-s < \epsilon$ .

Choose an integer  $n > \ell$  large enough so that

$$s \le \frac{i}{n} < \frac{i+1}{n} < \frac{i+2}{n} < \frac{i+3}{n} < s + \epsilon$$
 where  $i = \lfloor ns \rfloor + 1$ .

Then

$$B(\frac{j}{n}) - B(\frac{j-1}{n})| < \frac{7\ell}{n}$$
 for  $j = i+1, i+2, i+3$ .

### Proof.

Therefore the event that B(t) is differentiable at some point is contained in the set

$$B = \bigcup_{\ell \ge 1} \bigcup_{m \ge 1} \bigcap_{n \ge m} \bigcup_{0 \le i \le n+1} \bigcap_{i \le j \le i+3} \left\{ |B(\frac{j}{n}) - B(\frac{j-1}{n})| < \frac{7\ell}{n} \right\}.$$

We show P(B) = 0 as follows.

$$P\left(\bigcap_{n\geq m}\bigcup_{0\leq i\leq n+1}\bigcap_{1\leq j\leq i+3}\left\{|B(\frac{j}{n})-B(\frac{j-1}{n})|<\frac{7\ell}{n}\right\}\right)$$
  
$$\leq \liminf_{n\to\infty}P\left(\bigcup_{0\leq i\leq n+1}\bigcap_{1\leq j\leq i+3}\left\{|B(\frac{j}{n})-B(\frac{j-1}{n})|<\frac{7\ell}{n}\right\}\right)$$

Proof.

$$P\left(\bigcap_{n\geq m}\bigcup_{0\leq i\leq n+1}\bigcap_{i\leq j\leq i+3}\left\{|B(\frac{j}{n})-B(\frac{j-1}{n})|<\frac{7\ell}{n}\right\}\right)$$
  
$$\leq \liminf_{n\to\infty}P\left(\bigcup_{0\leq i\leq n+1}\bigcap_{i\leq j\leq i+3}\left\{|B(\frac{j}{n})-B(\frac{j-1}{n})|<\frac{7\ell}{n}\right\}\right)$$
  
$$\leq \liminf_{n\to\infty}\sum_{i=1}^{n+1}P\left(\bigcap_{i\leq j\leq i+3}\left\{|B(\frac{j}{n})-B(\frac{j-1}{n})|<\frac{7\ell}{n}\right\}\right)$$
  
$$\leq \liminf_{n\to\infty}n\left[P\left(|B(\frac{1}{n})|<\frac{7\ell}{n}\right)\right]^{3}$$
  
$$=\liminf_{n\to\infty}n\left[P\left(|B(1)|<\frac{7\ell}{\sqrt{n}}\right)\right]^{3}=\liminf_{n\to\infty}n\left[\frac{7\ell}{\sqrt{n}}\right]^{3}=0$$

**Stochastic Calculus** 

## Theorem Brownian motion is not Hölder of order 1/2

this follows from Modulus of continuity (P.Levy) With probability one,

$$\limsup_{\epsilon \to 0} \sup_{0 \le s < t \le 1 \atop t-s < \epsilon} \frac{|B_t - B_s|}{\sqrt{2\epsilon \log \epsilon^{-1}}} = 1$$

Lemma

$$\frac{x}{x^2+1}e^{-\frac{x^2}{2}} \le \int_x^\infty e^{-\frac{y^2}{2}}dy \le x^{-1}e^{-\frac{x^2}{2}} \qquad x > 0$$
$$\int_x^\infty e^{-\frac{y^2}{2}}dy \le x^{-1}\int_x^\infty ye^{-\frac{y^2}{2}}dy = x^{-1}e^{-\frac{x^2}{2}}$$



Proof of 
$$\limsup_{\epsilon \to 0} \sup_{\substack{0 \le s < t \le 1 \\ t-s < \epsilon}} \frac{|B_t - B_s|}{\sqrt{2\epsilon \log \epsilon^{-1}}} \ge 1$$

let 
$$\delta > 0$$
,  $A_n = \{\max_{1 \le k \le 2^n} |B_{\frac{k}{2^n}} - B_{\frac{k-1}{2^n}}| \le (1 - \delta)h(2^{-n})\},$   
 $h(t) = \sqrt{2t \log t^{-1}}$ 

$$P(A_n) \leq (1 - 2 \int_{(1-\delta)\sqrt{2\log 2^n}} \frac{e^{-y^2/2}}{\sqrt{2\pi}} dy)^{2^n} \quad independent \ increments$$
$$\leq e^{-Cn^{-1/2}2^{n(1-(1-\delta)^2)}} \quad by \ lemma$$

so  $\sum_{n=1}^{\infty} P(A_n) < \infty$ . By the Borel-Cantelli lemma, almost every  $\omega$  is in at most finitely many  $A_n$ .i.e.

$$\limsup_{\epsilon \to 0} \sup_{\substack{0 \le s < t \le 1\\ t-t-s < \epsilon}} \frac{|B_t - B_s|}{\sqrt{2\epsilon \log \epsilon^{-1}}} \ge 1 - \delta$$

now let  $\delta \downarrow 0$ 

Note: This proves that Brownian motion is not Hölder of order  $\alpha \ge 1/2$ 

Proof of 
$$\limsup_{\epsilon \to 0} \sup_{\substack{0 \le s < t \le 1 \\ t-s < \epsilon}} \frac{|B_t - B_s|}{\sqrt{2\epsilon \log \epsilon^{-1}}} \le 1$$

let  $\delta > 0$  and choose  $\epsilon > 0$  so that  $(1 + \epsilon)^2(1 - \delta) > 1 + \delta$  let

$$B_n = \{ \max_{i,j\in K} \frac{|B_{j/2^n} - B_{i/2^n}|}{h(k/2^n)} \ge 1 + \epsilon \}$$

$$K = \{ 0 \le i < j < 2^n, 0 < k = j - i \le 2^{n\delta} \}$$

$$\begin{array}{lcl} \mathcal{P}(B_n) & \leq & \sum_{K} \frac{2}{\sqrt{2\pi}} \int_{(1+\epsilon)\sqrt{\log(k^{-1}2^n)}}^{\infty} e^{-\frac{y^2}{2}} dy \\ & \leq & C \sum_{K} [\log(k^{-1}2^n)]^{-1/2} e^{-(1+\epsilon)^2 \log(k^{-1}2^n)} & \textit{lemma} \\ & \leq & C 2^{-n(1-\delta)(1-\epsilon)^2} \sum_{K} [\log(k^{-1}2^n)]^{-1/2} & k^{-1} \geq 2^{-n\delta} \end{array}$$

$$B_n = \{\max_{i,j\in K} \frac{|B_{j/2^n} - B_{i/2^n}|}{h(k/2^n)} \ge 1 + \epsilon\}$$

$$K = \{ 0 \le i < j < 2^n, 0 < k = j - i \le 2^{n\delta} \}$$

$$\begin{split} & P(B_n) \leq C 2^{-n(1-\delta)(1-\epsilon)^2} \sum_{K} [\log(k^{-1}2^n)]^{-1/2} \\ & \leq C n^{-1/2} 2^{n((1+\delta)-(1-\delta)(1+\epsilon)^2)} \quad |K| \leq 2^{n(1+\delta)}, \log(k^{-1}2^n) \geq \log 2^{n(1-\delta)} \end{split}$$

summable so with probability one there is an *N* s.t. for n > N, for  $i, j \in K$ ,

$$|B_{j/2^n} - B_{i/2^n}| < (1 + \epsilon)h(k/2^n)$$

Let  $\gamma > 0$ . Pick *N* large so that  $\sum_{m=n+1}^{\infty} h(2^{-m}) \leq \gamma h(2^{-(n+1)(1-\delta)})$ ,  $n \geq N$ Suppose that  $t = i2^{-n} + 2^{-n_1} + 2^{-n_2} + \cdots$  with  $N \leq n < n_1 < n_2 < \cdots$  and  $i \in K$ ,

$$|B_t - B_{i/2^n}| \le (1 + \epsilon) \sum_{m=n+1}^{\infty} h(2^{-m}) \le (1 + \epsilon)\gamma h(2^{-(n+1)(1-\delta)})$$

now suppose we have  $0 \le s < t \le 1$  and the special *n* so that

$$2^{-(n+1)(1-\delta)} \le t - s < 2^{-n(1-\delta)}$$

has  $n \ge N$  then we can write

$$\begin{aligned} |B_t - B_s| &\leq |B_{i/2^n} - B_s| + |B_{j/2^n} - B_{i/2^n}| + |B_t - B_{j/2^n}| \\ &\leq 2(1+\epsilon)\gamma h(2^{-(n+1)(1-\delta)}) + (1+\epsilon)h((j-i)2^{-n}) \\ &\leq (2(1+\epsilon)\gamma + 1 + \epsilon)h(t-s) \quad \text{if } t-s \text{ small enough} \end{aligned}$$

let  $\delta \downarrow 0$  and then  $\gamma \downarrow 0$